

# MTH 565

## Lectures 1 - 8

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**Topics:**

- Markov chains.
- Stationary distribution.
- Recurrent and transient states.

## Markov chains.

Consider a sequence of random variables  $X_0, X_1, X_2, \dots$  with values in a discrete **state space**  $S$ .

The sequence  $\{X_t\}_{t=0,1,\dots}$  is said to be a discrete time **Markov chain** if it satisfies the following property, known as **Markov property**:

$$P(X_{t+1} = j \mid X_t = i, X_{t-1} = i_{t-1}, \dots, X_1 = i_1, X_0 = i_0) = P(X_{t+1} = j \mid X_t = i)$$

A Markov chain  $\{X_t\}_{t=0,1,\dots}$  is said to be **time homogeneous** if

$$P(X_{t+1} = j \mid X_t = i) = p(i, j) \quad \text{for all } t = 0, 1, 2, \dots$$

**Markov chains.**

Consider a **time homogeneous** Markov chain  $X_0, X_1, X_2, \dots$  with a discrete state space  $S$  and transition probabilities

$$P(X_{t+1} = j \mid X_t = i) = p(i, j) \quad \text{for all } t = 0, 1, 2, \dots$$

Matrix (operator)  $P = \left( p(i, j) \right)_{i, j \in S}$  is called the **transition probability matrix (operator)**. Then

$$\sum_{j \in S} p(i, j) = 1 \quad \forall i \in S$$

**Example.** Consider  $S = \{0, 1\}$  (two states) and

$$P = \begin{pmatrix} 1 - p & p \\ q & 1 - q \end{pmatrix}$$

**Birth-and-death chain.** Consider state space

$$S = \{0, 1, 2, \dots\}$$

and a Markov chain  $\{X_t\}_{t=0,1,\dots}$  on  $S$  with transition probabilities

$p(i, i + 1) = p_i$ ,  $p(i, i - 1) = q_i$ , and  $p(i, i) = r_i$   
satisfying  $q_0 = 0$  and  $q_i + r_i + p_i = 1 \quad \forall i$

$$P = \begin{pmatrix} r_0 & p_0 & 0 & 0 & \dots \\ q_1 & r_1 & p_1 & 0 & \dots \\ 0 & q_2 & r_2 & p_2 & \dots \\ 0 & 0 & q_3 & r_3 & \dots \\ \vdots & \dots & \dots & \dots & \dots \end{pmatrix}$$

This is a Markov chain with only nearest neighbor transitions.

**Distribution of  $X_t$ .**

Consider a time homogeneous Markov chain  $\{X_t\}_{t=0,1,\dots}$  on a discrete state space  $S$ .

Define the **distribution** of  $X_t$  as a **row vector** of probability mass functions

$$\mu_t = \left( P(X_t = i) \right)_{i \in S}$$

ordered in the same way as enumeration of  $S$  used in  $P = \left( p(i, j) \right)_{i, j \in S}$ .

**Proposition.**

$$\mu_{t+1} = \mu_t P, \quad \text{and therefore} \quad \mu_t = \mu_0 P^t.$$

**Distribution of  $X_t$ .**

Define the **distribution** of  $X_t$  as a **row vector** of probability mass functions

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**Proposition.**

$$\mu_{t+1} = \mu_t P, \quad \text{and therefore} \quad \mu_t = \mu_0 P^t.$$

*Proof.* Denote by  $\mu_t(j)$  the  $j$ -th coordinate of  $\mu_t$ . Then

$$\mu_{t+1}(j) = P(X_{t+1} = j) = \sum_{i \in S} P(X_t = i) P(X_{t+1} = j | X_t = i) = \sum_{i \in S} \mu_t(i) p(i, j)$$

□

**Chapman-Kolmogorov equation.**

Consider a time homogeneous Markov chain  $\{X_t\}_{t=0,1,\dots}$  on a discrete state space  $S$ . For an integer  $s \geq 0$ , let

$$p_s(i, j) = P(X_{t+s} = j \mid X_t = i) \quad \forall t \geq 0$$

be the  $s$  step transition probabilities, and let

$$P_s = \left( p_s(i, j) \right)_{i, j \in S}$$

Observe that  $P_0 = I$  and  $P_1 = P$ .

**Chapman-Kolmogorov equation.**

$$P_{s+t} = P_s P_t \quad \forall s, t \geq 0, \quad \text{and therefore} \quad P_t = P^t$$

*Proof.*  $p_{s+t}(i, j) = \sum_{k \in S} p_s(i, k) p_t(k, j)$  follows from

$$P(X_{s+t} = j \mid X_0 = i) = \sum_{k \in S} P(X_s = k \mid X_0 = i) P(X_{s+t} = j \mid X_s = k) \quad \square$$

**Stationary distribution.**

For a homogeneous Markov chain with the transition probability matrix  $P = \left( p(i, j) \right)_{i, j \in S}$ , the **stationary distribution** (aka 'equilibrium distribution')  $\pi$  is defined as follows:

$$\pi P = \pi \quad \Leftrightarrow \quad \sum_{i \in S} \pi(i) p(i, j) = \pi(j).$$

Thus  $\sum_i \pi(i) p(i, j) = \pi(j) \sum_i p(j, i)$ , and for any state  $j \in S$ ,

$$\sum_{i: i \neq j} \pi(i) p(i, j) = \sum_{i: i \neq j} \pi(j) p(j, i).$$

Thus when restated in terms of traffic flow, the influx to the state  $j$  is equal to outflow from  $j$ , for each  $j$ . Thus the distribution stays unchanged.

### Stationary distribution.

Consider a homogeneous Markov chain over a discrete state space  $S$  with transition matrix  $P$ .

**Definition.** A homogeneous Markov chain is said to be **irreducible** if for any pair  $x, y \in S$ , there exists an integer  $k \geq 1$  such that  $p_k(x, y) > 0$ .

**Definition.** A homogeneous Markov chain is said to be **aperiodic** if

$$\gcd\{k \geq 1 : p_k(x, x) > 0\} = 1 \quad \forall x \in S$$

Notice that if the Markov chain is **irreducible**, then

$\gcd\{k \geq 1 : p_k(x, x) > 0\} = \gcd\{k \geq 1 : p_k(y, y) > 0\}$   
for any pair  $x, y \in S$ .

**Irreducible Markov chains.**

**Definition.** A homogeneous Markov chain is said to be **irreducible** if for any pair  $x, y \in S$ , there exists an integer  $k \geq 1$  such that  $p_k(x, y) > 0$ .

**Example (reducible).** Consider  $S = \{1, 2, 3, 4, 5\}$  and

$$P = \begin{pmatrix} 0.5 & 0.5 & 0 & 0 & 0 \\ 0.5 & 0.5 & 0 & 0 & 0 \\ 0 & 0 & 0.3 & 0.2 & 0.5 \\ 0 & 0 & 0.5 & 0.3 & 0.2 \\ 0 & 0 & 0.2 & 0.5 & 0.3 \end{pmatrix}$$

$\pi' = (1/2, 1/2, 0, 0, 0)$  and  $\pi'' = (0, 0, 1/3, 1/3, 1/3)$  are both stationary distributions, as well as all

$$\pi = \lambda\pi' + (1 - \lambda)\pi'' \quad \lambda \in [0, 1].$$

**Aperiodic Markov chains.**

**Definition.** A homogeneous Markov chain is said to be **aperiodic** if

$$\gcd\{k \geq 1 : p_k(x, x) > 0\} = 1 \quad \forall x \in S$$

**Example (periodic).** Consider  $S = \{0, 1\}$  (two states) and

$$P = \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix}$$

Here

$$\gcd\{k \geq 1 : p_k(x, x) > 0\} = 2 \quad \forall x \in S$$

The stationary distribution  $\pi = (0.5, 0.5)$  is unique.

However,  $\lim_{t \rightarrow \infty} p_t(x, y)$  does not exist.

**Stationary distribution.**

The following is version of ergodicity result for Markov chains over a **finite** state space  $S$ .

**Theorem (Ergodicity).** Consider an **irreducible** homogeneous Markov chain over a **finite** state space  $S$ . Then there exists a unique **stationary distribution**  $\pi$ . Furthermore, if the Markov chain is **aperiodic**, then

$$\lim_{t \rightarrow \infty} p_t(x, y) = \pi(y) \quad \forall x, y \in S.$$

In other words, for any distribution  $\mu_0$  of  $X_0$ ,

$$\mu_t = \mu_0 P^t \longrightarrow \pi \quad \text{as } t \rightarrow \infty.$$

**Example.** Consider  $S = \{0, 1\}$  (two states) and

$$P = \begin{pmatrix} 0.2 & 0.8 \\ 1 & 0 \end{pmatrix}$$

## Stationary distribution.

**Majorization proof of convergence.** Consider an **irreducible** and **aperiodic** homogeneous Markov chain over a **finite** state space  $S$ . Let  $\pi$  denote the unique **stationary distribution**.

We will prove that for any distribution  $\mu_0$  of  $X_0$ ,

$$\mu_t = \mu_0 P^t \longrightarrow \pi \quad \text{as } t \rightarrow \infty.$$

Without loss of generality let  $S = \{1, 2, \dots, d\}$ .

Can show that irreducibility and aperiodicity imply

$\exists m \in \mathbb{N}$  such that  $P^m > 0$ , i.e.  $P^m(i, j) > 0 \quad \forall i, j \in S$ .

Write

$$P^m = \begin{bmatrix} \text{---} & r_1 & \text{---} \\ \text{---} & r_2 & \text{---} \\ \vdots & \vdots & \vdots \\ \text{---} & r_d & \text{---} \end{bmatrix}, \quad \text{where row vectors } r_j > 0 \text{ (positive coord.)}$$

**Majorization proof of convergence (continued).**

$\exists m \in \mathbb{N}$  such that  $P^m > 0$ , i.e.  $P^m(i, j) > 0 \quad \forall i, j \in S$ .

Write

$$P^m = \begin{bmatrix} \text{---} & r_1 & \text{---} \\ \text{---} & r_2 & \text{---} \\ \vdots & \vdots & \vdots \\ \text{---} & r_d & \text{---} \end{bmatrix}, \quad \text{where row vectors } r_j > 0 \text{ (positive coord.)}$$

Thus, there exists  $\epsilon \in (0, 1)$  such that

$$\epsilon \pi \leq r_j \quad (\text{coordinatewise}) \quad \forall j \in S.$$

**Claim:** for any distribution  $\mu_0$  and  $k = 1, 2, \dots$ ,

$$\mu_{km} = (1 - (1 - \epsilon)^k) \pi + \nu_k,$$

where  $\nu_k \geq 0$  (coordinatewise) and  $\sum_{i=1}^d \nu_k(i) = (1 - \epsilon)^k$ .

**Majorization proof of convergence (continued).**

$\exists \epsilon \in (0, 1)$  s.t.  $\epsilon\pi \leq r_j$  (coordinatewise)  $\forall j \in S$ .

**Claim:** for any distribution  $\mu_0$  and  $k = 1, 2, \dots$ ,

$$\mu_{km} = (1 - (1 - \epsilon)^k)\pi + \nu_k,$$

where  $\nu_k \geq 0$  (coordinatewise) and  $\sum_{i=1}^d \nu_k(i) = (1 - \epsilon)^k$ .

Case  $k = 1$ :

$$\mu_m = \mu_0 P^m = \sum_{j=1}^d \mu_0(j) r_j = \epsilon\pi \sum_{j=1}^d \mu_0(j) + \sum_{j=1}^d \mu_0(j) (r_j - \epsilon\pi)$$

Hence,  $\mu_m = \epsilon\pi + \nu_1$ , where  $\nu_1 = \sum_{j=1}^d \mu_0(j) (r_j - \epsilon\pi)$

satisfies  $\nu_1 \geq 0$  (coordinatewise) and  $\sum_{i=1}^d \nu_1(i) = 1 - \epsilon$ .

Case  $k = 1$ :  $\mu_m = \mu_0 P^m = \epsilon\pi + \nu_1$ .

Induction step: Suppose  $\mu_{km} = (1 - (1 - \epsilon)^k)\pi + \nu_k$ , where  $\nu_k \geq 0$  (coordinatewise) and  $\sum_{i=1}^d \nu_k(i) = (1 - \epsilon)^k$ .

$\mu_{(k+1)m} = \mu_{km} P^m = (1 - (1 - \epsilon)^k)\pi + (1 - \epsilon)^k \mu'_0 P^m$ , where  $\mu'_0 = (1 - \epsilon)^{-k} \nu_k$  is a distribution.

By case  $k = 1$ , we have  $\mu'_0 P^m = \epsilon\pi + \nu'_1$ , where  $\nu'_1 \geq 0$  (coordinatewise) and  $\sum_{i=1}^d \nu'_1(i) = 1 - \epsilon$ . Thus,

$\mu_{(k+1)m} = (1 - (1 - \epsilon)^k)\pi + (1 - \epsilon)^k (\epsilon\pi + \nu'_1) = (1 - (1 - \epsilon)^{k+1})\pi + \nu_{k+1}$ , where  $\nu_{k+1} = (1 - \epsilon)^k \nu'_1$  satisfies

$\nu_{k+1} \geq 0$  and  $\sum_{i=1}^d \nu_{k+1}(i) = (1 - \epsilon)^k \sum_{i=1}^d \nu'_1(i) = (1 - \epsilon)^{k+1}$ .

We established the following claim.

**Claim:** for any distribution  $\mu_0$  and  $k = 1, 2, \dots$ ,

$$\mu_{km} = \mu_0 P^{km} = (1 - (1 - \epsilon)^k) \pi + \nu_k,$$

where  $\nu_k \geq 0$  (coordinatewise) and  $\sum_{i=1}^d \nu_k(i) = (1 - \epsilon)^k$ .

Thus, since  $\epsilon \in (0, 1)$ ,

$$\lim_{k \rightarrow \infty} \mu_0 P^{km} = \lim_{k \rightarrow \infty} \left[ (1 - (1 - \epsilon)^k) \pi + \nu_k \right] = \pi.$$

for any distribution  $\mu_0$ .

Hence, for any  $r = 0, 1, \dots, m - 1$ , we have

$$\lim_{k \rightarrow \infty} \mu_{km+r} = \lim_{k \rightarrow \infty} \mu_r P^{km} = \pi.$$

Therefore, we have proven that for any distribution  $\mu_0$ ,

$$\lim_{t \rightarrow \infty} \mu_t = \pi.$$

### Stationary distribution.

**Example.** Consider  $S = \{0, 1\}$  (two states) and

$$P = \begin{pmatrix} 1-p & p \\ q & 1-q \end{pmatrix} \quad p, q \in (0, 1)$$

**Eigenvalues:**  $\lambda_1 = 1$  and  $\lambda_2 = 1 - p - q$ .

Matrix theory:

For an eigenvalue  $\lambda_i$  of  $P$ , the **left eigenvector**  $u_i$  is a nonzero **row vector** satisfying  $u_i P = \lambda_i u_i$ .

The **right eigenvector**  $v_i$  is a nonzero **column vector** satisfying  $P v_i = \lambda_i v_i$ .

**Spectral Theorem:** for ‘reversible’ Markov chain, one can construct **biorthogonal** ( $u_i v_j = 0$  for  $i \neq j$ ) pair of left and right eigenbases  $\{u_i\}_{i \in S}$  and  $\{v_i\}_{i \in S}$ , and

$$P = \sum_i \lambda_i \frac{v_i u_i}{u_i v_i}, \quad \text{and therefore,} \quad P^t = \sum_i \lambda_i^t \frac{v_i u_i}{u_i v_i}$$

**Stationary distribution.**

**Example.** Consider  $S = \{0, 1\}$  (two states) and

$$P = \begin{pmatrix} 1-p & p \\ q & 1-q \end{pmatrix} \quad p, q \in (0, 1)$$

**Eigenvalues:**  $\lambda_1 = 1$  and  $\lambda_2 = 1 - p - q$ .

**Left eigenvectors:**  $u_1 = \pi = \left(\frac{q}{p+q}, \frac{p}{p+q}\right)$  and  $u_2 = (1, -1)$

**Right eigenvectors:**  $v_1 = \mathbf{1} = \begin{pmatrix} 1 \\ 1 \end{pmatrix}$  and  $v_2 = \begin{pmatrix} p \\ -q \end{pmatrix}$

**Spectral Theorem:**

$$\begin{aligned} P^t &= \sum_i \lambda_i^t \frac{v_i u_i}{u_i v_i} = \mathbf{1}\pi + (1-p-q)^t \frac{v_2 u_2}{p+q} \\ &= \begin{pmatrix} \frac{q}{p+q} & \frac{p}{p+q} \\ \frac{q}{p+q} & \frac{p}{p+q} \end{pmatrix} + \frac{(1-p-q)^t}{p+q} \begin{pmatrix} p & -p \\ -q & q \end{pmatrix} \end{aligned}$$

### Stationary distribution.

**Example.** Consider  $S = \{0, 1\}$  (two states) and

$$P = \begin{pmatrix} 1-p & p \\ q & 1-q \end{pmatrix} \quad p, q \in (0, 1)$$

Eigenvalues:  $\lambda_1 = 1$  and  $\lambda_2 = 1 - p - q$ .

Left eigenvectors:  $u_1 = \pi = \left( \frac{q}{p+q}, \frac{p}{p+q} \right)$  and  $u_2 = (1, -1)$

Right eigenvectors:  $v_1 = \mathbf{1} = \begin{pmatrix} 1 \\ 1 \end{pmatrix}$  and  $v_2 = \begin{pmatrix} p \\ -q \end{pmatrix}$

### Spectral Theorem:

$$P^t = \sum_i \lambda_i^t \frac{v_i u_i}{u_i v_i} = \mathbf{1}\pi + (1-p-q)^t \frac{v_2 u_2}{p+q} \longrightarrow \mathbf{1}\pi = \begin{pmatrix} \frac{q}{p+q} & \frac{p}{p+q} \\ \frac{q}{p+q} & \frac{p}{p+q} \end{pmatrix}$$

as  $t \rightarrow \infty$ . Hence,  $\lim_{t \rightarrow \infty} p_t(x, y) = \pi(y) \quad \forall x, y \in S$ .

## Stationary distribution.

**Theorem (Ergodicity).** Consider an **irreducible** homogeneous Markov chain over a **finite** state space  $S$ . Then there exists a unique **stationary distribution**  $\pi$ . Furthermore, if the Markov chain is **aperiodic**, then

$$\lim_{t \rightarrow \infty} p_t(x, y) = \pi(y) \quad \forall x, y \in S.$$

In other words, for any distribution  $\mu_0$  of  $X_0$ ,

$$\mu_t = \mu_0 P^t \longrightarrow \pi \quad \text{as } t \rightarrow \infty.$$

The **linear algebraic proof** uses the **Perron-Frobenius Theorem** to prove the above version of ergodicity for Markov chains with **finite** state space  $S$ .

**Perron-Frobenius Theorem.** If  $A = (a_{i,j}) \in \mathbb{R}^{d \times d}$  is an **irreducible nonnegative** matrix, then there exists a real eigenvalue  $\rho \geq 0$ , called **Perron-Frobenius eigenvalue**, such that

- $|\lambda| \leq \rho$  for any other eigenvalue  $\lambda$ .
- $\rho$  is **simple** (i.e., its left and right eigenspaces are one-dimensional), and it has left and right eigenvectors ( $u$  and  $v$ ) whose coordinates are **all positive**.

- $$\min_i \sum_{j=1}^d a_{i,j} \leq \rho \leq \max_i \sum_{j=1}^d a_{i,j}.$$

- $A$  has exactly  $h$  eigenvalues in  $\mathbb{C}$  with absolute value  $\rho$ :

$$\rho e^{2\pi i k/h} \quad k = 0, \dots, h-1,$$

where  $h$  is the **period** of  $A$ .

- $\lim_{n \rightarrow \infty} \frac{1}{\rho^n} A^n = \frac{vu}{uv}$ , where  $uA = \rho u$  and  $Av = \rho v$ .

**Stationary distribution.**

Matrix  $A$  is **irreducible** if  $\forall i, j \exists m \in \mathbb{N}$  such that  $(A^m)_{i,j} > 0$ .

Consider an **irreducible** homogeneous Markov chain over a **finite** state space  $S$ . Then  $P$  is an **irreducible nonnegative** matrix. Hence, Perron-Frobenius Theorem applies.

The **Perron-Frobenius eigenvalue**  $\rho$  of  $P$  satisfies

$$1 = \min_i \sum_{j=1}^d p(i, j) \leq \rho \leq \max_i \sum_{j=1}^d p(i, j) = 1.$$

Hence,  $\rho = 1$ .

Finally,  $\rho = 1$  is a **simple** eigenvalue, and it has left and right eigenvectors with **all positive** coordinates:

There exists a unique distribution  $\pi$  such that  $\pi P = \pi$ .

We know that  $P\mathbf{1} = \mathbf{1}$  as the rows of  $P$  add up to 1.

### Stationary distribution.

Consider an **irreducible** homogeneous Markov chain over a **finite** state space  $S$ . Then  $P$  is an **irreducible nonnegative** matrix. Hence, Perron-Frobenius Theorem applies with  $\rho = 1$ .

There exists a unique distribution  $\pi$  such that  $\pi P = \pi$ .

We know that  $P\mathbf{1} = \mathbf{1}$  as the rows of  $P$  add up to 1.

For an **irreducible** homogeneous Markov chain, matrix  $P$  has period

$$h = \gcd\{k \geq 1 : p_k(x, x) > 0\} \quad \text{same for all } x \in S.$$

By Perron-Frobenius Theorem,  $P$  has exactly  $h$  eigenvalues with absolute value  $\rho = 1$ :

$$e^{2\pi ik/h} \quad k = 0, \dots, h-1.$$

### Stationary distribution.

For an **irreducible** homogeneous Markov chain, matrix  $P$  has period

$$h = \gcd\{k \geq 1 : p_k(x, x) > 0\} \quad \text{same for all } x \in S.$$

By Perron-Frobenius Theorem,  $P$  has exactly  $h$  eigenvalues with absolute value  $\rho = 1$ :

$$e^{2\pi ik/h} \quad k = 0, \dots, h-1.$$

If  $P$  is aperiodic ( $h = 1$ ),

$$\lim_{t \rightarrow \infty} P^t = \frac{\mathbf{1}\pi}{\pi\mathbf{1}} = \mathbf{1}\pi.$$

If  $h > 1$ ,  $P^t$  does not have a limit as  $t \rightarrow \infty$ .

## Recurrent and transient states.

We will use the following notations:

$$P_x(A) = P(A | X_0 = x) \quad \text{and} \quad E_x[Y] = E[Y | X_0 = x].$$

For  $x \in S$ , consider the first hitting time

$$T_x = \min\{n \geq 1 : X_n = x\}.$$

**Definition.** A state  $x \in S$  is said to be **recurrent** if

$$P_x(T_x < \infty) = 1.$$

A recurrent state  $x \in S$  is **positive recurrent** if

$$E_x[T_x] < \infty.$$

Otherwise it is **null recurrent**.

**Definition.** A state  $x \in S$  is said to be **transient** if

$$P_x(T_x < \infty) < 1.$$

## Stationary distribution.

**Definition.** A state  $x \in S$  is said to be **positive recurrent** if

$$E_x[T_x] < \infty$$

The following is a version of ergodicity theorem for a general **discrete** state space  $S$ .

**Theorem (Ergodicity).** Consider an **irreducible** homogeneous Markov chain over a discrete state space  $S$ . If all of its states are **positive recurrent**, then there exists a unique **stationary distribution**  $\pi$  such that

$$\pi(x) = \frac{1}{E_x[T_x]}.$$

Furthermore, if the Markov chain is **aperiodic**,

$$\lim_{t \rightarrow \infty} p_t(x, y) = \pi(y) \quad \forall x, y \in S.$$